

INSTANTANEOUS FREQUENCY ESTIMATION USING A LEAST SQUARES TIME-FREQUENCY METHOD

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ABSTRACT

We present a method for estimating the instantaneous frequency of a signal. This method involves the calculation of a time-frequency energy density of the signal, then obtaining an instantaneous frequency estimation from this joint density. Time-frequency energy density is calculated as a least squares optimal combination of multi-window Gabor based evolutionary spectra. The optimal weights are obtained by minimizing an error criterion that is the difference between a reference time-frequency distribution and the combination of evolutionary spectra. Then instantaneous frequency of the signal is estimated from the final evolutionary spectrum as time conditional average frequency. Examples are given to illustrate the performance of our method.

Keywords: : Instantaneous frequency, Time-frequency analysis, Evolutionary spectrum.

1. INTRODUCTION

Instantaneous frequency (IF) of a signal, $\omega(t)$, is defined as the derivative of the phase of its corresponding analytic signal, $x(t) = A(t)e^{j\phi(t)}$ [1]. Moreover, from a joint time--frequency (TF) perspective, the IF of a signal is defined as the average of frequencies at a given time (or time conditional mean frequency) [2]:

$$\langle \omega \rangle_t = \omega(t) = \int \omega \frac{S(t, \omega)}{S(t)} d\omega \quad (1)$$

where

$$S(t) = \int S(t, \omega) d\omega$$

is the density in time ($|x(t)|^2$) or time marginal of the TF density $S(t, \omega)$. Estimating the IF of a signal is an important issue in many signal processing applications such as communications,

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radar, bioengineering, etc. [3,4]. For instance, in spread spectrum communication systems, jammers can be eliminated by estimating their IF and removing them by a time-varying filter [5].

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In our approach the IF is estimated using a least squares multi-window evolutionary spectrum as the TF energy density for the signal. TF signal analysis is a helpful tool for analyzing the time-varying frequency content of a non-stationary signal [2]. The Wigner-Ville Spectrum (WVS) is defined as a time-dependent spectrum for non-stationary stochastic process $x(t)$ and given by [6]:

$$P(t, \omega) = E\{W(t, \omega)\} \\ = E\left\{\int_{-\infty}^{\infty} \left[x\left(t - \frac{\tau}{2}\right)x^*\left(t + \frac{\tau}{2}\right)\right] e^{-j\omega\tau} d\tau\right\}$$

where $W(t, \omega)$ denotes the Wigner Distribution (WD) and the above is the statistical average of the WDs of the realizations of the process. When we have several observations of the non-stationary process $x(t)$, we can use an ensemble average of the individual WDs of these observations to estimate the WVS. However, this is not the case in general; we are only given a single realization of the process. In that case, Time-Frequency Distributions (TFDs) with a smoothing kernel function is used to estimate the WVS [2]. A good amount of research has been done to design kernels with desired properties yielding unbiased and low variance WVS estimates [6, 8].

A new estimate of the WVS is proposed as the optimal average of multiple-window spectrograms of the process in [9, 10]. In this work we use a WVS estimate that is an optimal combination of evolutionary spectra obtained by a multi-window Gabor expansion [7]. The optimal combination coefficients are obtained by minimizing the squared error between a reference TFD (which is taken to be the Wigner-Ville Distribution of the signal) and the multi-window spectral estimate.

2. EVOLUTIONARY SPECTRAL ANALYSIS BY MULTI-WINDOW GABOR EXPANSION

Given a non-stationary signal, $x(n), 0 \leq n \leq N-1$, a discrete Wold-Cramer representation [12] for it is given by

$$x(n) = \sum_{k=0}^{K-1} A(n, \omega_k) e^{j\omega_k n}, \quad (2)$$

where $\omega_k = 2\pi k/K$, K is the number of frequency samples, and $A(n, \omega_k)$ is an evolutionary kernel. The evolutionary spectrum is obtained from this kernel as

$$S(n, \omega_k) = \frac{1}{K} |A(n, \omega_k)|^2. \quad \text{In [7] we show}$$

that the kernel can be obtained from the coefficients of a Gabor expansion. The multi-window Gabor expansion is given by [7]

$$x(n) = \sum_{m=0}^{M-1} \sum_{k=0}^{K-1} a_{i,m,k} h_i(n - mL) e^{j\omega_k n} \quad (3)$$

$$= \sum_{k=0}^{K-1} A_i(n, \omega_k) e^{j\omega_k n} \quad (4)$$

where $\{a_{i,m,k}\}$ are the Gabor coefficients, $\{h_{i,m,k}\}$ are the Gabor basis functions that are obtained by scaling, translating and modulating with a sinusoid a window function:

$$h_{i,m,k}(n) = h_i(n - mL) e^{j\omega_k n} \quad (5)$$

and the synthesis window $h_i(n)$ is obtained by scaling a unit-energy mother window $g(n)$ as

$$h_i(n) = 2^{i/2} g(2^i n), i = 0, 1, \dots, I-1.$$

The multi-window Gabor coefficients are evaluated by

$$a_{i,m,k} = \sum_{n=0}^{N-1} x(n) \gamma_i^*(n - mL) e^{-j\omega_k n}, \quad (6)$$

where the analysis window $\gamma_i(n)$ is solved from the bi-orthogonality condition between $h_i(n)$ and $\gamma_i(n)$ [7]. Hence by comparing the representations of the signal in (3) and (4) we obtain the evolutionary kernel as

$$A_i(n, \omega_k) = \sum_{m=0}^{M-1} a_{i,m,k} h_i(n - mL) \quad (7)$$

Replacing for the coefficients $\{a_{i,m,k}\}$, we obtain also that

$$A_i(n, \omega_k) = \sum_{l=0}^{N-1} x(l) w_i(n, l) e^{-j\omega_k l}, \quad (8)$$

where we defined the time-varying window for scale i as

$$w_i(n, l) = \sum_{m=0}^{M-1} \gamma_i^*(l - mL) h_i(n - mL).$$

Then the evolutionary spectrum of $x(n)$ calculated using the window $h_i(n)$ is obtained by

$$S_i(n, \omega_k) = \frac{1}{K} |A_i(n, \omega_k)|^2,$$

where the factor $1/K$ is used for proper energy normalization. We should mention that normalizing the $w_i(n, l)$ to unit energy, the total energy of the signal is preserved thus justifying the use of $S_i(n, \omega_k)$ as a TF energy density for $x(n)$. Furthermore, $S_i(n, \omega_k)$ is always non-negative and approximates the marginal conditions [2]; hence, in contrast to many TFDs, interpretable as TF energy density function [7].

3. IF ESTIMATION BY MULTI-WINDOW LEAST SQUARES EVOLUTIONARY SPECTRUM

Given a realization of a discrete-time, nonstationary process corrupted by additive noise $x(n) = s(n) + \eta(n)$ where $s(n)$ and $\eta(n)$ denotes the signal and noise processes respectively. We intend to obtain a high resolution evolutionary spectral estimate with good performance in low signal to noise ratio (SNR) conditions such that the IF of the signal $s(n)$ can be estimated. We calculate a weighted

average combination of evolutionary spectra $S_i(n, \omega_k)$ that is closest to a reference TFD in a least squares sense. Given the signal $x(n)$, we calculate evolutionary spectra $S_i(n, \omega_k)$ for $i = 0, 1, \dots, I-1$ as

$$S_i(n, \omega_k) = \frac{1}{K} \left| \sum_{l=0}^{N-1} x(l) W_i(n, l) e^{-j\omega_k l} \right|^2. \quad (9)$$

Gauss windows are used as $h_i(n)$, for their optimal concentration in the TF plane [7]. Then we estimate the WVS of the process $x(n)$ as a weighted average of the evolutionary spectra

$$\hat{P}(n, \omega_k) = \sum_{i=0}^{I-1} c_i S_i(n, \omega_k) \quad (10)$$

where the weights $\{c_i\}$ are obtained by minimizing the error function

$$\mathcal{E}_i = \sum_{n=0}^{N-1} \sum_{k=0}^{K-1} \left| P_R(n, \omega_k) - \sum_{i=0}^{I-1} c_i S_i(n, \omega_k) \right|^2 \quad (11)$$

and $P_R(n, \omega_k)$ is a reference TFD which is taken here as Wigner-Ville Distribution of the signal for its optimal TF resolution. By using a matrix notation, the minimization problem in (11) can be rewritten as

$$\min_{c_i} \|P_R - S c\|^2 \quad (12)$$

The solution of this least squares minimization problem is

$$c^o = (S^T S)^{-1} S^T P_R$$

where the superscript ' o ' stands for optimum. Then a WVS estimate is obtained as optimal weighted average using $\{c_i^o\}$ in equation (10). Finally, we mask or threshold our estimate $\hat{P}(n, \omega_k)$ to eliminate any possible negative values caused by any negative c_i^o coefficient,

and result in a non-negative time-varying spectrum, i.e.,

$$\hat{P}(n, \omega_k)^+ = \begin{cases} \hat{P}(n, \omega_k), & \hat{P}(n, \omega_k) \geq 0; \\ 0, & \hat{P}(n, \omega_k) < 0. \end{cases} \quad (13)$$

where $\hat{P}(n, \omega_k)^+$ denotes the positive-only part of the spectrum. Then the IF of the signal can be calculated from this TF density according to equation (1) as time conditional mean frequency. In our simulations we obtain both TF density and IF estimate of several nonstationary signals.

4. EXAMPLES

To illustrate the performance of our proposed method, we consider a quadratic FM signal. In Fig.1, we show the least squares evolutionary spectral estimate $\hat{P}(n, \omega_k)^+$ of this signal. Fig. 2 shows the IF estimate (the solid line) obtained from this evolutionary spectrum, and the true IF of the signal (the dashed line). As shown the IF estimate of the signal is very close to the true IF. To compare our result with that of another method, we applied Cakrak and Loughlin's [10] optimal combination of spectrograms method and obtained the result given in Fig. 3. As another example, we consider a sinusoidal FM signal and estimate its IF using our method. Fig. 4 shows the least squares evolutionary spectral estimate of this signal. The IF estimate (the solid line), and the true IF of the signal (the dashed line) are given in Fig. 5. IF estimate obtained by the least squares combination of spectrograms is shown in Fig. 6. It is clear from the figures that our least squares combination of multi-window evolutionary spectra gives better IF estimates than Cakrak and Loughlin's [10] spectrogram method.

5. CONCLUSIONS

In this work, we present a new method for obtaining the Instantaneous Frequency of non-stationary signals. Our method uses the optimal combination in the least squares sense, of evolutionary spectra that are calculated by multi-window Gabor expansion. The optimal weights are obtained by minimizing the squared error between the combination of evolutionary spectra and a reference TFD. Examples show that our method combines the advantages of multiple-

window evolutionary spectral analysis and high resolution TFDs, i.e., it provides non-negative and high resolution time-varying spectral estimates as such the IF estimate is sufficiently close to the correct IF of the signal.

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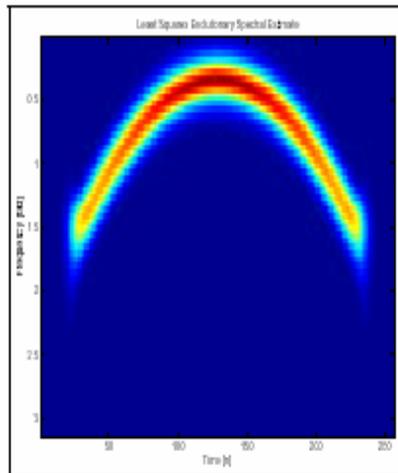


Fig. 1. Least Squares Multi-window Evolutionary Spectral Estimate.

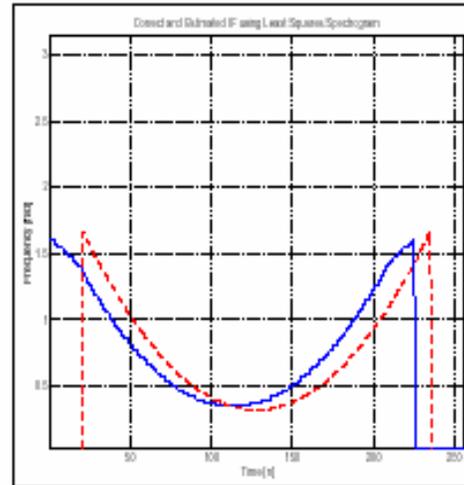


Fig. 3. Correct (dashed line) and estimated (solid line) IF by spectrogram combinations.

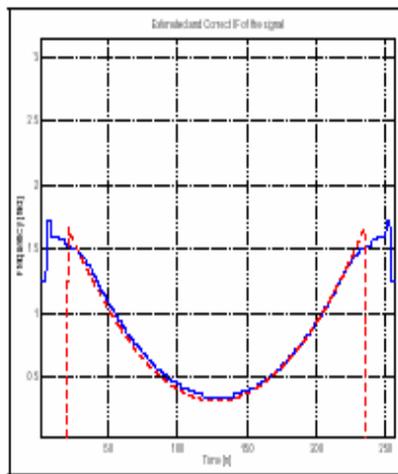


Fig. 2. Estimated (solid line) and true (dashed line) IF.

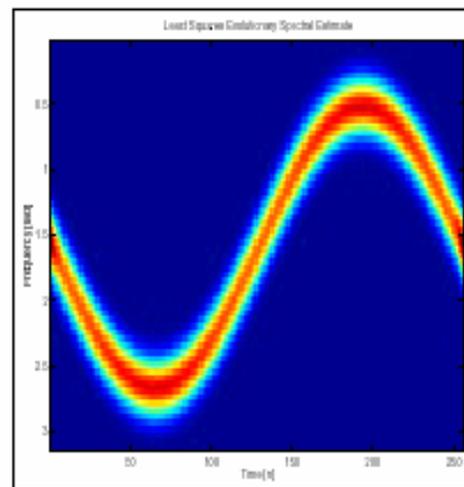


Fig. 4. Least Squares Evolutionary Spectral Estimate of the signal.

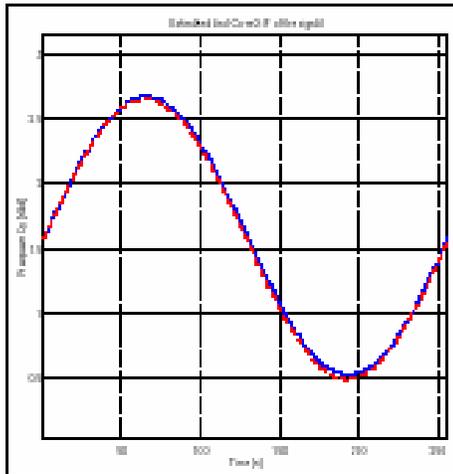


Fig. 5. Estimated (solid line) and true (dashed line) IF.

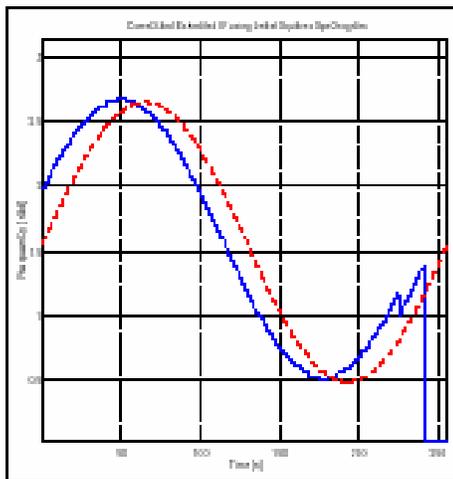


Fig. 6. Correct (dashed line) and estimated (solid line) IF by spectrogram combinations.

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